

# Complementary proofs to "Approximation of discrete BSDE using least-squares regression"

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## Abstract

This short note gives complementary proofs to our work [GT11], of which we follow the notations and assumptions.

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## 1. Assumption (A<sub>F</sub>-iii) for non-uniform grids

The time grid  $(t_k = T - T(1 - k/N)^{1/\theta_\pi})_{0 \leq k \leq N}$  with  $\theta_\pi \in (0, 1]$  satisfies

$$C_\pi := \sup_{k < N} \frac{\Delta_k}{(T - t_k)^{1-\theta_L}} \leq \frac{T^{\theta_L}}{\theta_\pi} \frac{1}{N^{1 \wedge \frac{\theta_L}{\theta_\pi}}},$$

$$R_\pi := \sup_{0 \leq k \leq N-2} \frac{\Delta_k}{\Delta_{k+1}} \leq \frac{1}{\theta_\pi} \left( 1 \vee \left( \frac{1}{2\theta_\pi} \right)^{\frac{1}{\theta_\pi}-1} \right),$$

where  $\Delta_k = t_{k+1} - t_k$  and  $\theta_L \in (0, 1]$ .

PROOF. Set  $1/\theta_\pi = \mu \geq 1$  and  $g(x) = 1 - (1 - x)^\mu$ : we have  $t_k = Tg(k/N)$ . Note that  $g$  is increasing and concave; thus we have

$$\frac{\Delta_k}{(T - t_k)^{1-\theta_L}} \leq \frac{\frac{\mu T}{N}(1 - k/N)^{\mu-1}}{T^{1-\theta_L}(1 - k/N)^{\mu(1-\theta_L)}} = \frac{T^{\theta_L}}{\theta_\pi N} (1 - k/N)^{\theta_L/\theta_\pi-1}$$

and the bound on  $C_\pi$  follows by considering either  $\theta_L \geq \theta_\pi$  or  $\theta_L < \theta_\pi$ .

Now, we study  $R_\pi$ . Since  $g$  is concave, we have  $\Delta_{k-1} \geq \Delta_k \geq \dots \geq \Delta_{N-1} = TN^{-\mu}$  and  $\Delta_{k-1} \leq \frac{\mu T}{N}(1 - k/N)^{\mu-1}$ . This gives a first upper bound for the

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$n_0$ -last times  $k = N - n_0, \dots, N - 1$  (with  $n_0 \geq 1$ ):

$$\frac{\Delta_{k-1}}{\Delta_k} \leq \frac{\Delta_{k-1}}{\Delta_{N-1}} \leq \frac{\frac{\mu T}{N}(1 - k/N)^{\mu-1}}{TN^{-\mu}} = \mu(N - k)^{\mu-1} \leq \mu n_0^{\mu-1}. \quad (1)$$

We are now in a position to complete the upper bound on  $R_\pi$ .

- $\mu \in [1, 2]$ : we prove  $\frac{\Delta_{k-1}}{\Delta_k} \leq \mu$ . For  $k = N - 1$ , the inequality is true owing to (1). Now take  $k < N - 1$ . Since  $g''$  is non-increasing ( $\mu \in [1, 2]$ ), we have  $\Delta_k \geq \frac{\mu T}{N}(1 - k/N)^{\mu-1} + \frac{T}{2N^2}g''((k+1)/N)$ , and we easily deduce

$$\frac{\Delta_{k-1}}{\Delta_k} \leq \frac{\frac{\mu T}{N}(1 - k/N)^{\mu-1}}{\frac{\mu T}{N}(1 - k/N)^{\mu-1} - \frac{T}{2N^2}\mu(\mu - 1)(1 - (k+1)/N)^{\mu-2}} \leq \frac{1}{1 - \frac{(\mu-1)}{2}} \leq \mu.$$

- $\mu \geq 2$ : we prove  $\frac{\Delta_{k-1}}{\Delta_k} \leq \mu(\frac{\mu}{2})^{\mu-1}$ . Set  $n_0 = \lfloor \frac{\mu}{2} \rfloor$ :  $n_0 \leq \frac{\mu}{2} < n_0 + 1$ . For  $k \geq N - n_0$ , the announced upper bound directly follows from (1). Now take  $k \leq N - n_0 - 1$  (which implies  $N - k > \frac{\mu}{2}$ ):  $g''$  being non-decreasing for  $\mu \geq 2$ , we have

$$\begin{aligned} \Delta_k &\geq \frac{\mu T}{N}(1 - k/N)^{\mu-1} + \frac{T}{2N^2}g''(k/N) \\ &= \frac{\mu T}{N}(1 - k/N)^{\mu-1} \left[ 1 - \frac{(\mu-1)}{2}(N - k)^{-1} \right] \\ &> \frac{\mu T}{N}(1 - k/N)^{\mu-1} \frac{1}{\mu} \geq \Delta_{k-1} \frac{1}{\mu}. \end{aligned}$$

□

## 2. Proof of Proposition 3.3.

**Proposition 3.3.** Assume  $(\mathbf{A}'_\xi\text{-i})$  and  $(\mathbf{A}_F)$ . For any  $R \in [0, +\infty]$  and for any  $\pi$  with  $N$  large enough (such that  $C_\pi L_f^2 \leq \frac{1}{12q}$ ), the following almost sure error bounds on  $Y_i - Y_i^R$  and  $Z_i - Z_i^R$  hold for any  $0 \leq i < N$ :

$$\begin{aligned} |Y_i - Y_i^R| &\leq C_y \exp\left(\frac{T}{8} + \frac{12qL_f^2}{\theta_L}T^{\theta_L}\right) \exp\left(-\frac{1}{4}R^2\right)\sqrt{N}, \\ \left(\sum_{k=i}^{N-1} \mathbb{E}_i |Z_k - Z_k^R|^2 \Delta_k\right)^{\frac{1}{2}} &\leq C_y \exp\left(\frac{12qL_f^2}{\theta_L}T^{\theta_L}\right) (8q + T \exp(\frac{T}{4}))^{\frac{1}{2}} \exp\left(-\frac{1}{4}R^2\right)\sqrt{N}. \end{aligned}$$

PROOF. We set  $\mathcal{T}_R := \mathbb{E}([\mathcal{N} - (-R) \vee \mathcal{N} \wedge R]^2)$  where  $\mathcal{N}$  is a Gaussian random variable with mean 0 and variance 1. An explicit computation gives

$$\mathcal{T}_R = 2(\mathbb{P}(\mathcal{N} > R)(R^2 + 1) - R \frac{e^{-\frac{1}{2}R^2}}{\sqrt{2\pi}}) \leq 2\mathbb{P}(\mathcal{N} > R)(R^2 + 1 - R^2) \leq 2e^{-\frac{1}{2}R^2},$$

where the two last inequalities are derived from the Mill inequality and the Markov exponential inequality.

Now, we follow the arguments of Lemma 3.1 and we consider  $\gamma \in (0, +\infty)^N$  such that  $8q(\Delta_k + \frac{1}{\gamma_k}) \frac{L_f^2}{(T-t_k)^{1-\theta_L}} \leq 1$  for  $0 \leq k < N$ . Define  $\Delta Y_k := Y_k - Y_k^R$  and  $\Delta Z_k := Z_k - Z_k^R$ .

*Preliminary bound on  $\Delta Z$ .* Applying the Cauchy-Schwartz inequality and the almost sure bounds on  $Y$  and  $Y^R$  (Proposition 3.2), we obtain:

$$\begin{aligned} \Delta_k |\Delta Z_k|^2 &= \Delta_k^{-1} |\mathbb{E}_k[Y_{k+1} \Delta W_k - Y_{k+1}^R [\Delta W_k]_w]|^2 \\ &\leq 2\Delta_k^{-1} |\mathbb{E}_k[Y_{k+1}(\Delta W_k - [\Delta W_k]_w)]|^2 + 2\Delta_k^{-1} |\mathbb{E}_k[\Delta Y_{k+1} [\Delta W_k]_w]|^2 \\ &\leq 2qC_y^2 \mathcal{T}_R + 2q(\mathbb{E}_k[\Delta Y_{k+1}^2] - (\mathbb{E}_k[\Delta Y_{k+1}])^2). \end{aligned} \quad (2)$$

*Bound on  $\Delta Y$ .* Using Young's inequality  $(a+b)^2 \leq (1+\Delta_k \gamma_k) a^2 + (1+\frac{1}{\Delta_k \gamma_k}) b^2$ , the Lipschitz property of  $(y, z) \mapsto f_k(y, z)$ , and using (2), we obtain

$$\begin{aligned} \Delta Y_k^2 &\leq (1 + \Delta_k \gamma_k) (\mathbb{E}_k[\Delta Y_{k+1}])^2 \\ &\quad + 2(\Delta_k + \frac{1}{\gamma_k}) \frac{L_f^2}{(T-t_k)^{1-\theta_L}} \Delta_k (\mathbb{E}_k[\Delta Y_{k+1}^2] + |\Delta Z_k|^2) \end{aligned} \quad (3)$$

$$\begin{aligned} &\leq (1 + \Delta_k \gamma_k - 4q(\Delta_k + \frac{1}{\gamma_k}) \frac{L_f^2}{(T-t_k)^{1-\theta_L}}) (\mathbb{E}_k[\Delta Y_{k+1}])^2 \\ &\quad + 2(\Delta_k + \frac{1}{\gamma_k}) \frac{L_f^2}{(T-t_k)^{1-\theta_L}} (\Delta_k + 2q) \mathbb{E}_k[\Delta Y_{k+1}^2] \\ &\quad + 4qC_y^2 (\Delta_k + \frac{1}{\gamma_k}) \frac{L_f^2}{(T-t_k)^{1-\theta_L}} \mathcal{T}_R. \end{aligned} \quad (4)$$

The condition  $8q(\Delta_k + \frac{1}{\gamma_k}) \frac{L_f^2}{(T-t_k)^{1-\theta_L}} \leq 1$  ensures that  $1 + \Delta_k \gamma_k - 4q(\Delta_k + \frac{1}{\gamma_k}) \frac{L_f^2}{(T-t_k)^{1-\theta_L}} \geq 0$ ; this given, we may use Jensen's inequality on the term (4) to obtain:

$$\begin{aligned} \Delta Y_k^2 &\leq (1 + \Delta_k \gamma_k + 2(\Delta_k + \frac{1}{\gamma_k}) \frac{L_f^2}{(T-t_k)^{1-\theta_L}} \Delta_k) \mathbb{E}_k[\Delta Y_{k+1}^2] \\ &\quad + 4qC_y^2 (\Delta_k + \frac{1}{\gamma_k}) \frac{L_f^2}{(T-t_k)^{1-\theta_L}} \mathcal{T}_R \\ &\leq (1 + \Delta_k \gamma_k + \frac{\Delta_k}{4}) \mathbb{E}_k[\Delta Y_{k+1}^2] + \frac{1}{2} C_y^2 \mathcal{T}_R \end{aligned}$$

using again the relation between  $\Delta_k$  and  $1/\gamma_k$ . Multiplying by  $\lambda_k := \prod_{j=0}^{k-1} (1 + \Delta_j \gamma_j + \frac{\Delta_j}{4})$ , taking conditional expectation  $\mathbb{E}_i$ , and summing over  $k = i, \dots, N-1$ , we obtain a pointwise uniform bound for  $\Delta Y_i^2$ :

$$\Delta Y_i^2 \Gamma_i \leq \Delta Y_i^2 \lambda_i \leq \frac{1}{2} C_y^2 e^{T/4} \Gamma_N N \mathcal{T}_R. \quad (5)$$

Final bound on  $\Delta Z$ . (2) yields:

$$\begin{aligned} \sum_{k=i}^{N-1} \Gamma_k \mathbb{E}_i[|\Delta Z_k|^2] \Delta_k &\leq 2qC_y^2 \Gamma_N N \mathcal{T}_R + 2q \sum_{k=i}^{N-1} (\mathbb{E}_i[\Delta Y_{k+1}^2] - \mathbb{E}_i(\mathbb{E}_k[\Delta Y_{k+1}])^2) \Gamma_{k+1} \\ &\leq 2qC_y^2 \Gamma_N N \mathcal{T}_R + 2q \sum_{k=i}^{N-1} (\mathbb{E}_i[\Delta Y_k^2] - (1 + \Delta_k \gamma_k) \mathbb{E}_i(\mathbb{E}_k[\Delta Y_{k+1}])^2) \Gamma_k. \end{aligned}$$

Substituting the inequality (3), we obtain

$$\begin{aligned} \sum_{k=i}^{N-1} \Gamma_k \mathbb{E}_i[|\Delta Z_k|^2] \Delta_k &\leq 2qC_y^2 \Gamma_N N \mathcal{T}_R + 4q \sum_{k=i}^{N-1} (\Delta_k + \frac{1}{\gamma_k}) \frac{L_f^2}{(T - t_k)^{1-\theta_L}} \Delta_k \Gamma_k (\mathbb{E}_i[\Delta Y_k^2] + \mathbb{E}_i[|\Delta Z_k|^2]) \\ &\leq 2qC_y^2 \Gamma_N N \mathcal{T}_R + \frac{1}{2} \sum_{k=i}^{N-1} \Delta_k \Gamma_k (\mathbb{E}_i[\Delta Y_k^2] + \mathbb{E}_i[|\Delta Z_k|^2]) \end{aligned}$$

taking into account the relation between  $\pi$  and  $\gamma$ . Thus, we have

$$\sum_{k=i}^{N-1} \Gamma_k \mathbb{E}_i[|\Delta Z_k|^2] \Delta_k \leq C_y^2 \Gamma_N N \mathcal{T}_R (4q + \frac{T}{2} \exp(\frac{T}{4})). \quad (6)$$

Observe that  $\gamma_k := \frac{24qL_f^2}{(T-t_k)^{1-\theta_L}}$  defines an admissible choice, provided that  $C_\pi L_f^2 \leq \frac{1}{12q}$ . It gives  $1 \leq \Gamma_i \leq \Gamma_N \leq \exp(\frac{24qL_f^2}{\theta_L} T^{\theta_L})$ . Plugging this estimate into (5) and (6), and using the bound on  $\mathcal{T}_R$ , we obtain the final result.  $\square$

### 3. Proof of Lemma 4.2

**Lemma 4.2.** With the current notation and assumptions, for all  $m$  we have

$$\begin{aligned} \bar{y}_k^{R,M}(X_k^m) &= \mathbb{E}_N^M [\Phi(\tilde{X}_N^{k,m}) + \sum_{i=k}^{N-1} f_i(\tilde{X}_i^{k,m}, y_{i+1}^{R,M}(\tilde{X}_{i+1}^{k,m}), z_i^{R,M}(\tilde{X}_i^{k,m})) \Delta_i], \\ \Delta_k \bar{z}_{l,k}^{R,M}(X_k^m) &= \mathbb{E}_N^M [[\Delta \tilde{W}_{l,k}^m]_w(\Phi(\tilde{X}_N^{k,m}) + \sum_{i=k+1}^{N-1} f_i(\tilde{X}_i^{k,m}, y_{i+1}^{R,M}(\tilde{X}_{i+1}^{k,m}), z_i^{R,M}(\tilde{X}_i^{k,m})) \Delta_i)]. \end{aligned}$$

PROOF. We start with a standard result. Let  $\mathcal{G}$  and  $\mathcal{H}$  be sub- $\sigma$ -algebras of  $\mathcal{F}$ , such that  $\mathcal{G} \perp \mathcal{H}$ . Let  $F : \Omega \times \mathbb{R}^d \rightarrow \mathbb{R}^d$  be bounded and  $\mathcal{G} \otimes \mathcal{B}(\mathbb{R}^d)$ -measurable, and  $U : \Omega \rightarrow \mathbb{R}^d$  be  $\mathcal{H}$ -measurable. Then, by the Monotone Class Theorem for

functions,  $\mathbb{E}[F(U)|\mathcal{H}] = j(U)$  where  $j(h) = \mathbb{E}[F(h)]$  for all  $h \in \mathbb{R}^d$ .

In order to apply the above result, we require some standard results about the ghost path  $(\tilde{X}, \Delta\tilde{W})$ . Let  $k$  be fixed. Since  $\tilde{X}$  is a Markov chain, then for all  $i > k$  there is a mapping  $V_i : \Omega \times \mathbb{R}^d \rightarrow \mathbb{R}^d$  measurable with respect to  $\mathcal{G}_i \otimes \mathcal{B}(\mathbb{R}^d)$  such that  $\tilde{X}_i^{x,k} = V_i(x)$ , where the filtration  $(\mathcal{G}_i)_{k < i \leq N}$  is independent of  $\mathcal{F}_N^M$  and  $\Delta\tilde{W}_k^k$  is  $\mathcal{G}_{k+1}$ -measurable.

Now, by defining

$$\begin{cases} F_1(x) &:= \Psi_k^{R,M}(x, V_{k+1}(x), \dots, V_N(x)), \\ F_2(x) &:= [\Delta\tilde{W}_k^{k,m}]_w \Psi_{k+1}^{R,M}(V_{k+1}(x), V_{k+2}(x), \dots, V_N(x)), \end{cases}$$

the result of the previous paragraph can be applied, because  $F_1$  and  $F_2$  are  $\mathcal{G}_N \otimes \mathcal{B}(\mathbb{R}^d)$ -measurable, hence the representations for  $y_k^{R,M}(X_k^m)$  and  $\tilde{z}_k^{R,M}(X_k^m)$ .  $\square$

#### 4. Proof of Lemma 1 in Appendix B

**Lemma 1.** Let  $\mathcal{G}$  be a countable set of functions  $g : \mathbb{R}^d \mapsto [0, B]$  with  $B > 0$ . Let  $X, X^1, \dots, X^M$  ( $M \geq 1$ ) be i.i.d.  $\mathbb{R}^d$  valued random variables. For any  $\alpha > 0$  and  $\varepsilon \in (0, 1)$  one has

$$\begin{aligned} \mathbb{P}\left(\sup_{g \in \mathcal{G}} \frac{\frac{1}{M} \sum_{m=1}^M g(X^m) - \mathbb{E}[g(X)]}{\alpha + \frac{1}{M} \sum_{m=1}^M g(X^m) + \mathbb{E}[g(X)]} > \varepsilon\right) &\leq 4\mathbb{E}(\mathcal{N}_1(\frac{\alpha\varepsilon}{5}, \mathcal{G}, X^{1:M})) \exp\left(-\frac{3\varepsilon^2\alpha M}{40B}\right), \\ \mathbb{P}\left(\sup_{g \in \mathcal{G}} \frac{\mathbb{E}[g(X)] - \frac{1}{M} \sum_{m=1}^M g(X^m)}{\alpha + \frac{1}{M} \sum_{m=1}^M g(X^m) + \mathbb{E}[g(X)]} > \varepsilon\right) &\leq 4\mathbb{E}(\mathcal{N}_1(\frac{\alpha\varepsilon}{8}, \mathcal{G}, X^{1:M})) \exp\left(-\frac{6\varepsilon^2\alpha M}{169B}\right). \end{aligned}$$

PROOF. The first inequality is stated in [GKKW02, Theorem 11.6] for  $B \geq 1$ . For  $B \in (0, 1)$ , we rescale the class of functions  $\{g/B : g \in \mathcal{G}\}$  (now bounded by 1), replace  $\alpha$  by  $\alpha/B$  and apply the previous case: this gives the announced upper bound.

To establish the second inequality, we adapt the proof of the first inequality from the proof of [GKKW02, Theorem 11.6]. The first step consists in taking a ghost sample  $\tilde{X}^{1:M}$  and observing that for a given  $g \in \mathcal{G}$ ,  $\mathbb{E}[g(X)] - \frac{1}{M} \sum_{m=1}^M g(X^m) > \varepsilon(\alpha + \frac{1}{M} \sum_{m=1}^M g(X^m) + \mathbb{E}[g(X)])$  and  $\mathbb{E}[g(X)] - \frac{1}{M} \sum_{m=1}^M g(\tilde{X}^m) \leq \frac{\varepsilon}{4}(\alpha + \frac{1}{M} \sum_{m=1}^M g(\tilde{X}^m) + \mathbb{E}[g(X)])$  imply

$$\begin{aligned} (1 + \frac{5\varepsilon}{8}) &\left(\frac{1}{M} \sum_{m=1}^M g(\tilde{X}^m) - \frac{1}{M} \sum_{m=1}^M g(X^m)\right) \\ &> \frac{3\varepsilon}{8}(2\alpha + \frac{1}{M} \sum_{m=1}^M g(X^m) + \frac{1}{M} \sum_{m=1}^M g(\tilde{X}^m)) + \frac{3\varepsilon}{4}\mathbb{E}[g(X)]. \end{aligned}$$

Since the r.h.s. positive, the l.h.s. is also positive; using  $\frac{13}{8} \geq 1 + \frac{5\varepsilon}{8}$  implies

$$\frac{1}{M} \sum_{m=1}^M g(\tilde{X}^m) - \frac{1}{M} \sum_{m=1}^M g(X^m) > \frac{3\varepsilon}{13} (2\alpha + \frac{1}{M} \sum_{m=1}^M g(X^m) + \frac{1}{M} \sum_{m=1}^M g(\tilde{X}^m)).$$

Then we proceed as in [GKKW02, pp. 205-207] to show that the probability to estimate is bounded by

$$2\mathbb{P}\left(\exists g \in \mathcal{G} : \frac{1}{M} \sum_{m=1}^M g(\tilde{X}^m) - \frac{1}{M} \sum_{m=1}^M g(X^m) > \frac{3\varepsilon}{13} (2\alpha + \frac{1}{M} \sum_{m=1}^M g(X^m) + \frac{1}{M} \sum_{m=1}^M g(\tilde{X}^m))\right)$$

for  $M > \frac{8B}{\varepsilon^2\alpha}$  (however for  $M \leq \frac{8B}{\varepsilon^2\alpha}$  the upper bound in Lemma 1 is obviously true). The rest of the proof is identical to [GKKW02, pp. 208-210], except that one should take a  $\mathbf{L}_1$   $\delta$ -cover of  $\mathcal{G}$  w.r.t.  $X^{1:M}$  with  $\delta = \frac{\alpha\varepsilon}{8}$  (instead of  $\delta = \frac{\alpha\varepsilon}{5}$ ). It leads to a new upper bound,  $4\mathbb{E}(\mathcal{N}_1(\frac{\alpha\varepsilon}{8}, \mathcal{G}, X^{1:M})) \exp(-\frac{6\varepsilon^2\alpha M}{169B})$ .  $\square$

## References

- [GKKW02] L. Györfi, M. Kohler, A. Krzyżak, and H. Walk. *A distribution-free theory of nonparametric regression*. Springer Series in Statistics, 2002.
- [GT11] E. Gobet and P. Turkedjiev. Approximation of discrete BSDE using least-squares regression. *Submitted*, 2011.